

Sparse Representations and Bayesian Denoising

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Outline

- Introduction.
 - Observation model.
 - Bayesian paradigm.
- Ingredients from modern harmonic analysis.
- Statistical sparse priors:
 - Univariate (marginal).
 - Multivariate (joint).
- Bayesian estimation.
- Combining transforms.
- Conclusion and extensions.

Observation model

- The image (or signal) is viewed as realization(s) of a RV or a random field whose degradation equation is:

$$Y_s = \mathcal{M} [\Psi((\mathcal{B}X)_s) \odot \epsilon_s] \quad (1)$$

where:

- \odot is any composition of two arguments (e.g. '+', '·').
- $s \in \mathcal{S}$ is the location index.
- ϵ_s is the noise (random) (generally assumed AWGN but not necessarily so, e.g. speckle, Poisson, $\frac{1}{f}$).
- \mathcal{B} is a (possibly non-linear) degradation operator (e.g. convolution with a PSF).
- Ψ is a transformation not necessarily linear nor invertible (e.g. sensor-specific, etc).
- \mathcal{M} missing data mechanism.

Restoration problem: how to estimate unobserved X from observed Y

An inverse ill-posed problem

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Bayesian paradigm

$$Y = \Psi((B(X))) \odot \epsilon$$

• $p(x, \mathbf{z})$: *prior* distribution. \mathbf{z} some other image features (e.g. local regularity, texture, etc).

• $p(y|x, \mathbf{z})$: likelihood (given x and \mathbf{z}). ($p(\epsilon)$).

• $p(y)$: marginal distribution = $\int p(y|x, \mathbf{z})p(x, \mathbf{z})dx d\mathbf{z}$.

• $p(x, \mathbf{z}|y)$, *posterior* distribution: $\frac{p(y|x, \mathbf{z})p(x, \mathbf{z})}{p(y)}$

Bayesian estimation amounts to finding the operator \mathcal{D} s.t.:

$$\hat{x} = \arg \inf_{\mathcal{D} \in \mathcal{O}_n} R(x, \hat{x} = \mathcal{D}y) = \mathbb{E}_{Y, X} [L(x, \mathcal{D}y)] \quad (2)$$

What estimator for what risk ?

| Cost $L(x, \hat{x})$ | Estimator | \hat{x} |
|--|-----------|--|
| 0-1 | MAP | $\arg \max_{x \in \mathcal{X}} p_{X Y}(x y)$ |
| L_2 | MF | $E[X Y]$ |
| L_1 | MMAE | $\Pr(X > \hat{x} Y = y) = \frac{1}{2}$ |
| $\text{Card} \{s \in S : \hat{x}_s \neq x_s\}$ | MPM | MAP at each site. |

- MAP involves solving an optimization problem.
- MF involves solving an integration problem.
- For mutually independent iid gaussian signal and noise, MAP, MF and Wiener are the same.

What prior ?

● Image corpus based models

- Existence of a probability space on some particular corpus of images (e.g. natural images) [Olshausen and Field 96, Zhu and Mumford 96, Gousseau 00, Gousseau-Morel-Alvarez 99, Mumford and Huang 99].

● Transported Generator Models

- Random objects randomly placed (according to some probability law) while imposing some axioms such as scale invariance (multi-scale nature of images) [Gidas et Mumford 01, Grenander et al. 99-03].

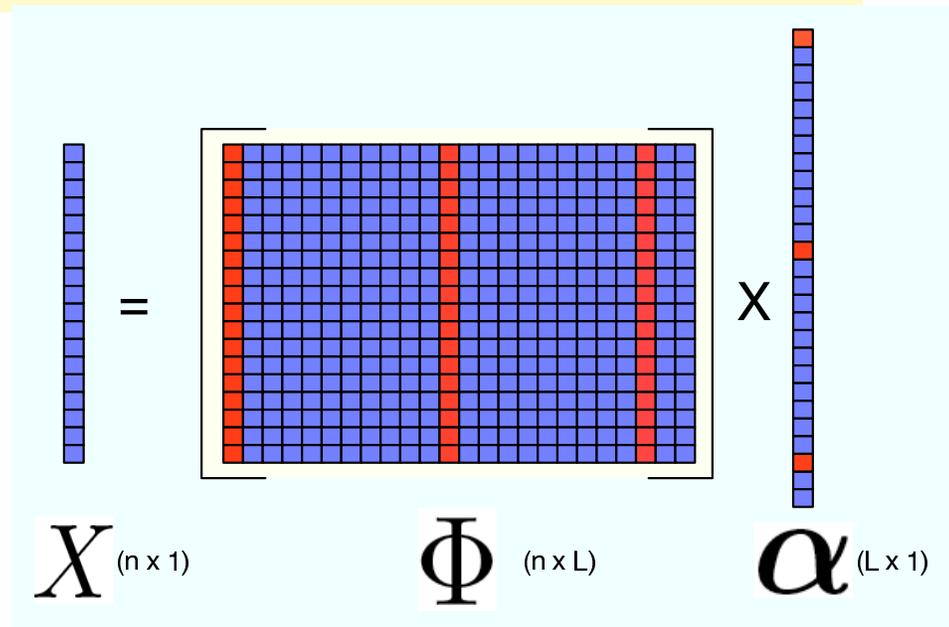
● Random Field Theory (e.g. MRF) [Besag 86, Geman and Geman 88].

● Sparse representation-based prior [Mallat 89, Simoncelli et al. 98, Wainwright et al. 00, Grenander et al. 01, Achim et al. 01, Fadili et al. 03, etc].

Elements from modern harmonic analysis

- An n -sample signal or image X can be written as the superposition of elementary functions $\phi_\gamma(s)$ (*atoms*) parameterized by γ s.t. (Γ is denumerable):

$$X_s = \sum_{\gamma \in \Gamma} \alpha_\gamma \phi_\gamma(s), \quad \phi_\gamma \in \mathcal{L}, \quad \|\alpha\|_0 \ll n$$



- The atoms $\{\phi_l\}_{l=1, \dots, L}$ are normalized to a unit ℓ_2 norm.
- $\Phi = [\phi_1 \dots \phi_L] \in \mathbb{R}^{N \times L}$, $\text{Card } \Gamma = L \geq N$ (bases, tight frames or frames).
- Examples of Γ : frequency (Fourier), scale-translation (wavelets), scale-translation-frequency (wavelet packets), translation-duration-frequency (cosine packets), scale-translation-angle (geometrical X-lets, curvelets, bandlets, contourlets, wedgelets, etc).

Sparse representation-based denoising

- Stable recovery of a sparse representation in presence of (AWG) noise (with bounded variance):

$$\min_{\alpha} \|\alpha\|_0 \quad \text{s.t.} \quad \|Y - \Phi\alpha\|_2 \leq \delta$$

- Convexify and relax [Chen et al. 01, Donoho et al. 04, Gribonval et al. 04, Fuchs 05, etc]:

$$\min_{\alpha} \frac{1}{2} \|Y - \Phi\alpha\|_2^2 + \lambda \|\alpha\|_1$$

- Corresponds to the MAP estimate with Laplacian prior.
- Here, the ℓ_1 norm prior will be replaced by a flexible family of sparsity promoting priors and the MAP by the MF.

A brief overview

Classical term-by-term

Minimax estimation, SureShrink, etc [Donoho et al. 92-95]. Modifications on Donoho's shrinkage operators [Bruce and Gao, Antoniadis and Fan]. Translation invariant threshold [Coifman and Donoho 95]. Hypothesis testing [Abramovich and Benjamini 95-96, Ogden and Parzen 96]. Cross-validation [Green et Silverman 94, Eubank 99].

Bayesian term-by-term (univariate)

Bernoulli-Gaussian FM [Abramovich et al. 98, Clyde and George 99,00]. Bayesian hypothesis testing [Vidakovic et al. 98]. SMG with exponential multiplier prior [Vidakovic et al. 00]. Two Gaussians FM [Chipman et al. 97]. t -Student prior [Vidakovic 98]. GGD [Mallat99, Liu et Moulin 99]. Adaptive variance gaussian prior [Simoncelli 99]. α -stable [Achim et al. 01].

Classical block

Non-overlapping block thresholding [Cai 99]. Overlapping block thresholding [Cai et Silverman 00].

Bayesian block (multivariate)

Non-overlapping block bayesian estimation [Abramovich et Sapatinas 00]. Multivariate gaussian prior [Huang and Cressie 00]. Mixed effects models [Huang and Lu 00]. MRF [Malfait et al. 97, Crouse et al. 98, Pizurica et al. 02]. HMT model [DSP Rice (Romberg, Baraniuk et al. 00-02)]. Scale mixture of gaussians [Li and Orchard 00, Mihchak et al. 99, Portilla et al. 03]. Multi-variate α -stable [Koruglu et Achim 04].

Univariate Scale Mixture of Gaussians family

Definition 1 (Andrews and Mallows 74) Let X be a RV with real-valued realizations. Under the SMG, there exist two independent RVs $U \geq 0$ and $Z \sim \mathcal{N}(0, 1)$ such that:

$$X \stackrel{d}{=} Z\sqrt{U} \quad (3)$$

Property 1

• SMG is a subset of the elliptically symmetric distributions [Kotz et al. 89]

• $f_X(0)$ exists iff $\mathbb{E}[U^{-1/2}] < +\infty$.

• The pdf of X is:

$$f_X(x) = \frac{1}{\sqrt{2\pi}} \int_0^{+\infty} u^{-1/2} e^{-\frac{x^2}{2u}} f_U(u) \quad (4)$$

It is unimodal, symmetric around the mode and differentiable almost everywhere (except perhaps at 0).

• The characteristic function (CF) of X is:

$$\Phi_X(\omega) = \mathcal{L}[f_U] \left(\frac{\omega^2}{2} \right) \quad (5)$$

\mathcal{L} is the Laplace transform.

• The pdf of U is closely related to the inverse Laplace transform of f_X .

USMG family

The following proposition establishes necessary and sufficient conditions for such a representation to exist:

Proposition 1 *The RV X has a SMG representation iff the k^{th} derivatives of $f_X(\sqrt{y})$ have alternating sign, i.e.:*

$$\left(-\frac{d}{dy}\right)^k f_X(\sqrt{y}) \geq 0 \quad \forall y > 0 \quad (6)$$

Lemma 1 *If $X \stackrel{d}{=} Z\sqrt{U}$ with random $U \geq 0$ and $Z \sim \mathcal{N}(0, \sigma^2)$, then $\text{kurtosis}(X) > 0 \implies$ the symmetric distribution of X is necessarily sharply peaked (leptokurtic) with heavy tails.*

- For sparse representations, empirical coefficient pdfs are symmetric around 0, leptokurtic and heavy tailed.
- These pdfs have their 1st and 2nd derivatives of alternating signs on \mathbb{R}^+ .
- The SMG family satisfies all these requirements.
- Consequence: this family is well adapted to capture the sparsity of decompositions and is then legitimate as a prior for the coefficients.
- A key advantage of SMG is that it transfers desirable properties of the gaussian distribution through the mixing RV.

Relating USMG to Besov space

- An explicit relationship between the parameters of the USMG prior on the wavelet coefficients of g and the Besov space within which g may fall (a.s.).

Theorem 1 *Let $X_{j,k} \stackrel{d}{=} Z_j \sqrt{U}$ iid RVs at each scale such that $Z \sim \mathcal{N}(0, \sigma_j^2)$, $E[U] = 1$ and $M_U(p) < +\infty$ $1 \leq p < +\infty$ and $\sigma_j = \sigma_0 2^{-j\beta}$ (the scale invariance property of images), with $(0 < \sigma_0 < +\infty, \beta \geq 0)$. Then, for a fixed $c_{0,0}$,*

$g \in B_{p,q}^s$ almost surely if and only if $\beta > (s + \frac{1}{2})$, for $1 \leq p < \infty$ and $1 \leq q \leq \infty$.

USMG: empirical Bayes

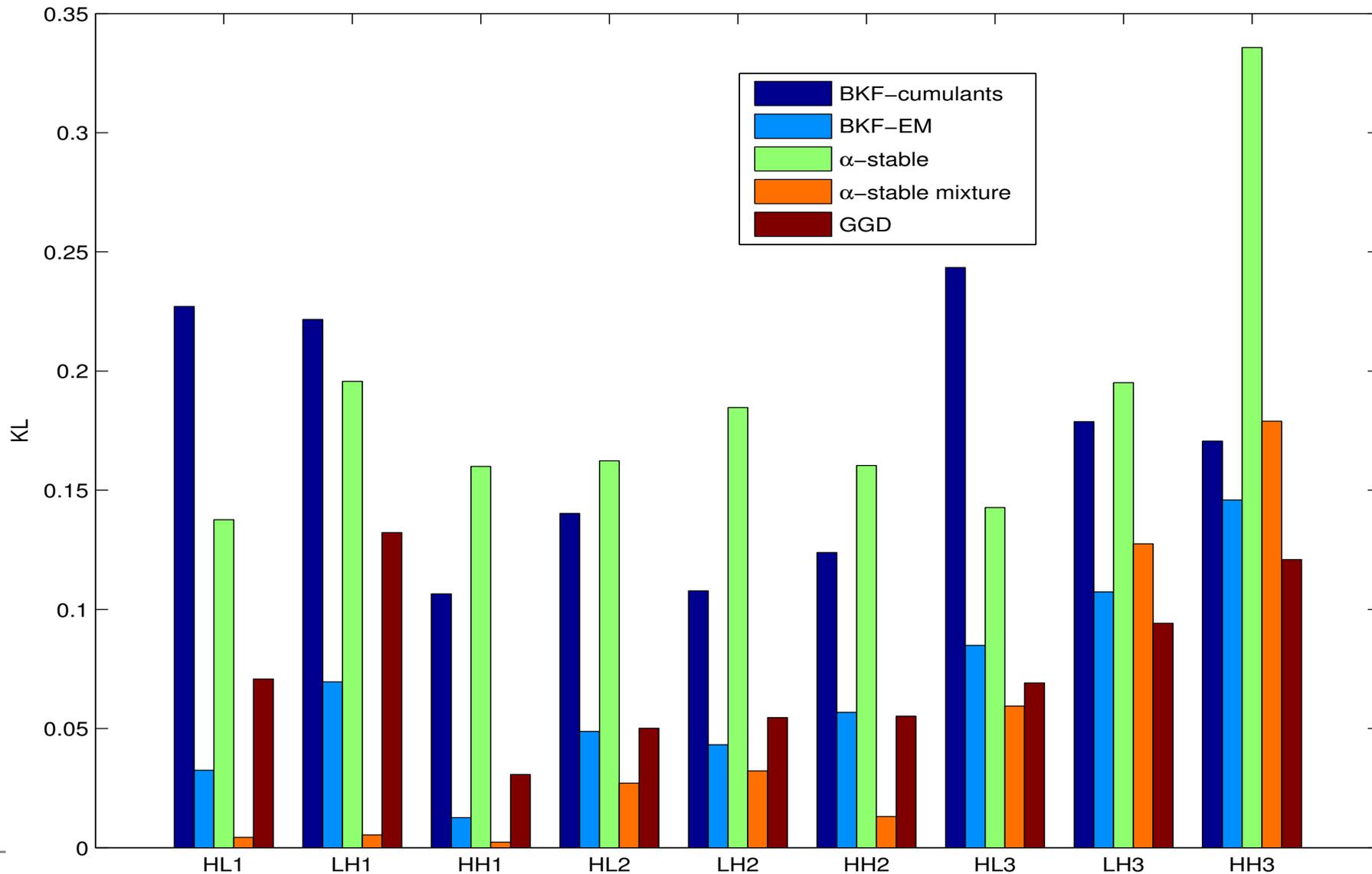
- Generally, distribution of U depends on some hyperparameters θ .
- They are estimated directly from the coefficients at each subband:
 - MLE.
 - Quantile methods.
 - Characteristic function methods.
 - **Cumulants** (easily extended in presence of AWGN).
 - **EM**.
- This step is crucial for the final performance of the denoiser.
- Is somewhat easy in noiseless case and becomes much more complex with corrupting noise.

USMG: examples

- The Bessel K form prior [Grenander et al. 03, Fadili et al. 03]: the mixing RV is Gamma distributed.
- The α -stable prior [Achim et al. 01, Boubchir and Fadili 03]: the mixing RV is also α -stable.
- The GGD prior [Mallat 89, Moulin and Liu 99]: the mixing RV is not known in a closed form.

Example on the DWT

Comparison on a 100 image database.



Multivariate SMG family

Definition 2 Let \mathbf{X} be a VRV taking values in \mathbb{R}^d . Under the SMG, there exists a RV $U \geq 0$ and a VRV $\mathbf{Z} \sim \mathcal{N}(0, \Sigma)$, $\Sigma > 0$ (U and \mathbf{Z} mutually independent) such that:

$$\mathbf{X} \stackrel{d}{=} \mathbf{Z}\sqrt{U} \quad (7)$$

Property 2

• MSMG is a sub-family of elliptically symmetric distributions [Kotz et al. 89].

• The pdf of \mathbf{X} is:

$$f_{\mathbf{X}}(x) = (2\pi)^{-d/2} |\Sigma|^{-1} \int_0^{+\infty} u^{-1/2} \exp \left[-\frac{\mathbf{x}^T \Sigma^{-1} \mathbf{x}}{2u} \right] f_U(u) du \quad (8)$$

• It is unimodal, elliptically symmetric with elliptically symmetric CF:

$$\Phi_{\mathbf{X}}(\omega) = \mathcal{L} [f_U] \left(\frac{\omega^T \Sigma \omega}{2} \right) \quad (9)$$

Multivariate SMG family

Lemma 2 *For a RV U ,*

- The measure of multivariate kurtosis of \mathbf{X} in the sense of Mardia is always strictly positive.*
 - A necessary and sufficient condition for a MSMG representation to exist is the alternation of sign of the derivatives of its functional parameter (density generator).*
- The multivariate SMG family satisfies the requirements of leptokurticity, heavy-tailness and symmetry.
 - This family is again adapted to capture the sparsity and dependency structure of the representation coefficients and is then legitimate as a multivariate prior.

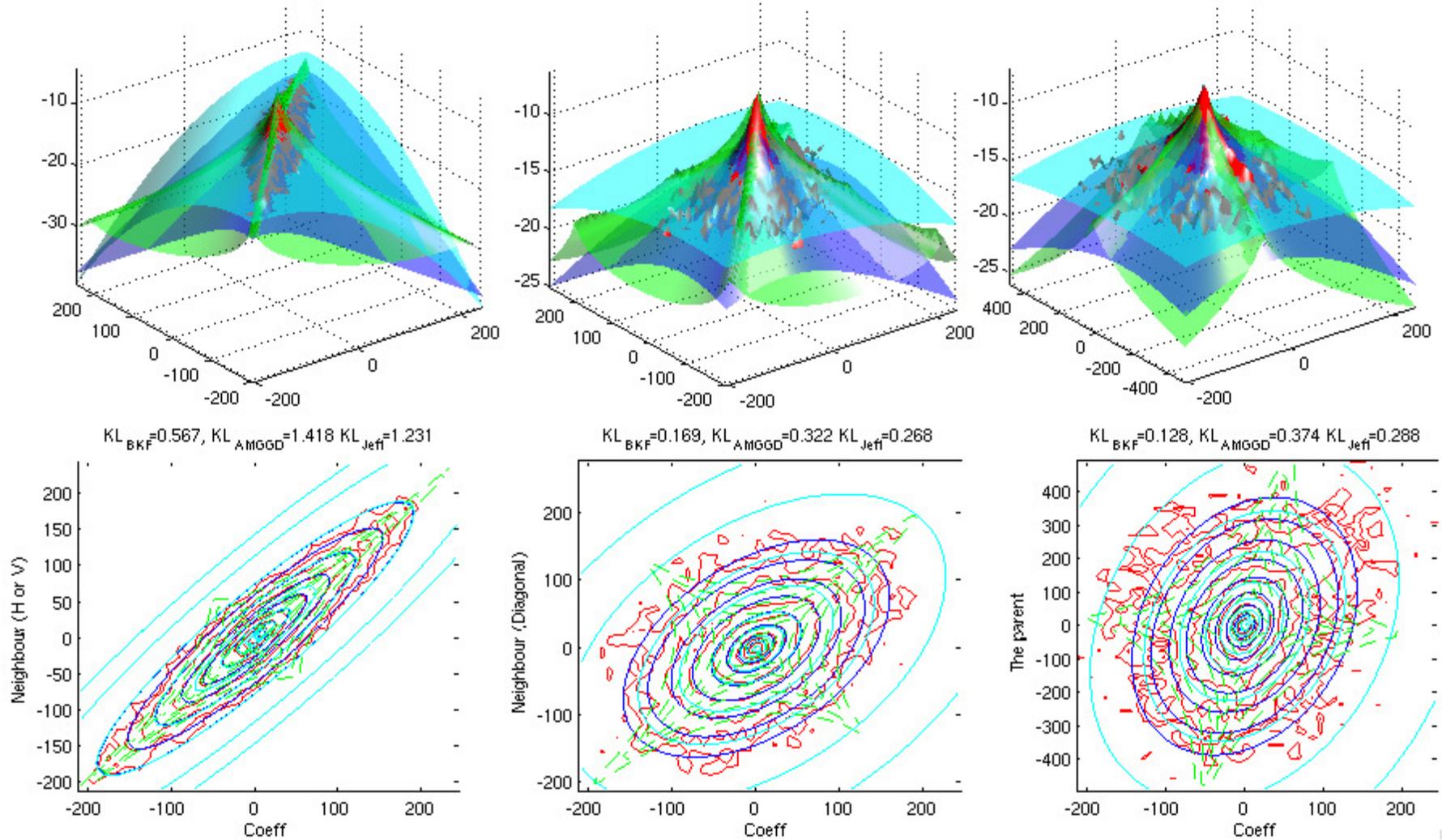
MSMG: empirical Bayes

- Again, the distribution of the multiplier U depends on some hyperparameters θ .
- Σ and these hyperparameters are estimated directly from the coefficients:
 - MLE (very time consuming).
 - **Moments and Cumulants** ($\mathbb{E} \left[(\mathbf{X}\Sigma^{-1}\mathbf{X})^i \right] = 2^i \frac{\Gamma(d/2+i)}{\Gamma(d/2)} \mathbb{E} [U^i]$).
 - **EM** (easily adapted if univariate EM is accessible).
- Again, this step is a chief obstacle towards good performance of the denoiser.
- Somewhat easy in noiseless case but more complex with corrupting noise.

MSMG: examples

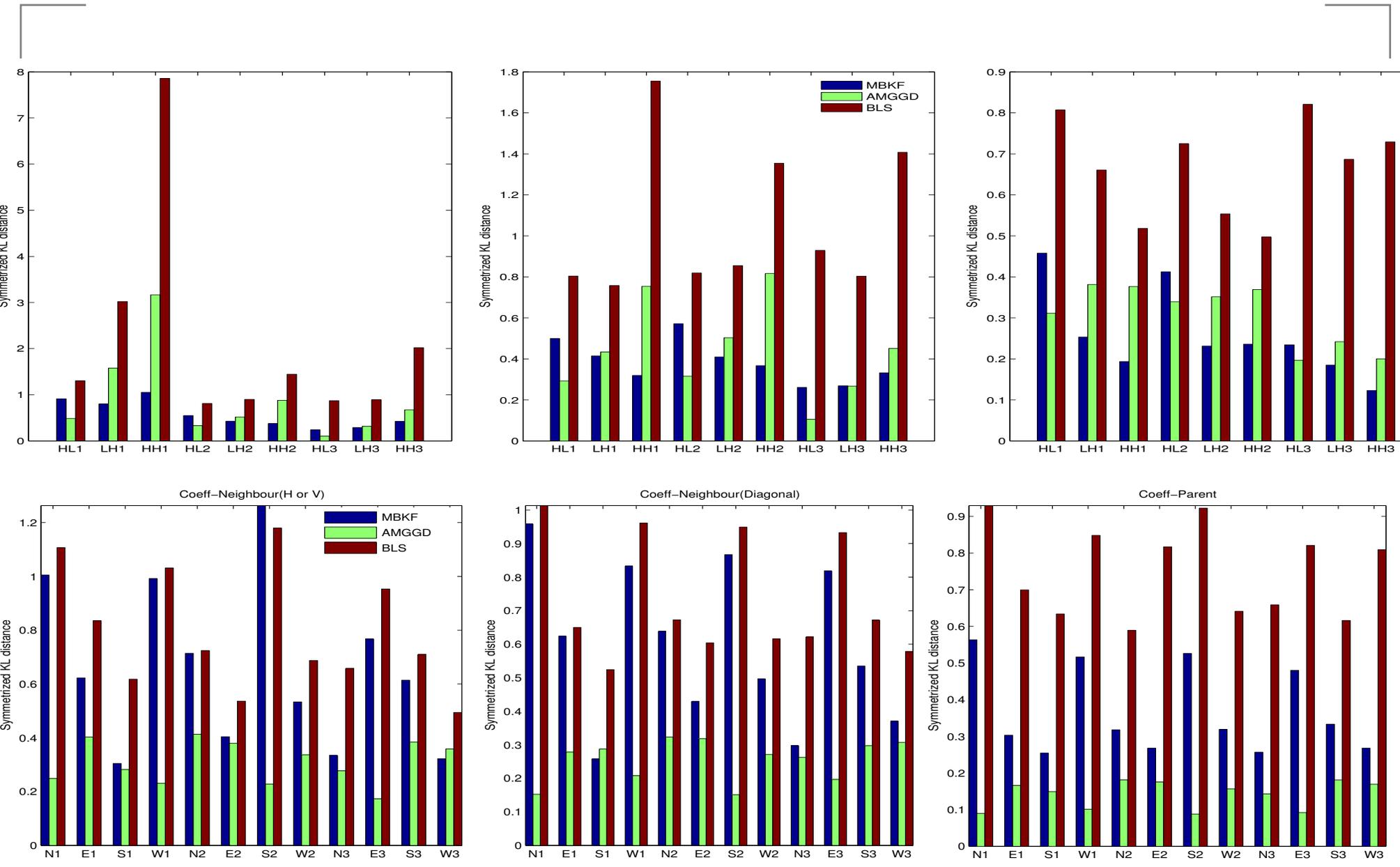
- The multivariate Bessel K form prior [Fadili et al. 06]: the mixing RV is Gamma distributed.
- The α -stable prior [Kuruoglu et al. 04]: the mixing RV is also α -stable.
- The GGD prior [Fadili et al. 05].
- Non-informative Jeffrey's prior [Portilla et al. 03].

Visual illustration ($d = 2$)



Empirical (red), MBKF (blue), AMGGD (green), MSMG with Jeffrey's multiplier (cyan).

Example on the DWT and FDCT



Application to Bayesian denoising

- The MF estimator corresponds to a scale mixture of local Wiener estimates:

$$\hat{\alpha}_{\text{MF}} = \frac{\int_0^{+\infty} \underbrace{\hat{\alpha}|u}_{\text{Wiener estimate}} \phi(\mathbf{d}; u\Sigma + \Sigma_\varepsilon) f_U(u) du}{\int_0^{+\infty} \phi(\mathbf{d}; u\Sigma + \Sigma_\varepsilon) f_U(u) du}$$

- Closed-form expressions.
- Deploy the bayesian integration technology:
 - Analytic approximation (e.g. Laplace, Saddlepoint).
 - Quadrature numerical integration (accurate but slow).
 - Monte-Carlo Integration (fast and accurate).

Application to Bayesian denoising (cont'd)

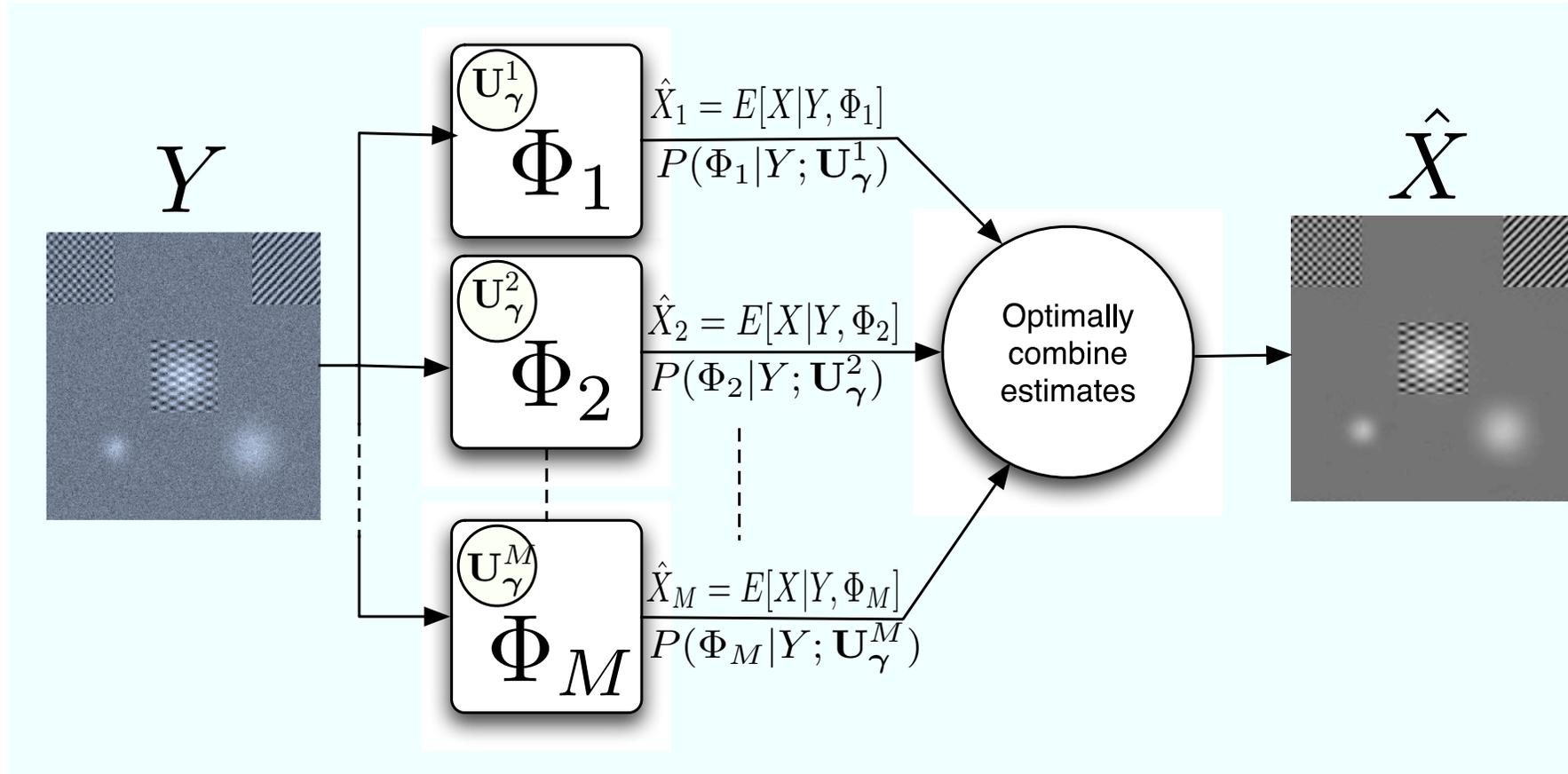
Main results

Theorem 2

- *Both the MF and the MAP estimators under the UBKF prior have closed analytical forms.*
- *The UBKF MAP estimator is equivalent to universal soft thresholding for $\frac{\sigma_\varepsilon^2}{c} = \log n$ as $\alpha \rightarrow 1$ (Laplacian prior) or large n .*
- *Bayesian CLT: the UBKF MAP estimator is asymptotically gaussian (as $n \rightarrow +\infty$).*
- *The MF estimator under the α -stable prior has a closed analytical form.*

Combining transforms: bases

- The dictionary is a union of M (sufficiently incoherent) bases $\{\Phi_m\}_{m=1,\dots,M}$.



$$\hat{X} = \sum_{m=1}^M \hat{X}_m P(\Phi_m|Y; U_\gamma^m)$$

Combining transforms: bases

- Suppose that:
 - The SMG prior is independent of the transform.
 - The transforms are equiprobable (or given by a learning step).
- For the USMG:

$$P(\Phi_m | Y_s; \mathbf{U}_\gamma^m) = \frac{\int \cdots \int_0^{+\infty} \mathcal{N}(Y_s; \underbrace{\mu_m}_{\text{LP component}}, V_m(s) + \sigma_\varepsilon^2) \underbrace{f_{U_\gamma^m}(u_\gamma^m)}_{\text{SMG prior}} du_1^m \cdots du_{|\Gamma_m|}^m}{\sum_{m'=1}^M \int \cdots \int_0^{+\infty} \mathcal{N}(Y_s; \mu_{m'}, V_{m'}(s) + \sigma_\varepsilon^2) f_{U_{\gamma'}^{m'}}(u_{\gamma'}^{m'}) du_1^{m'} \cdots du_{|\Gamma_{m'}|}^{m'}}$$

$$V_m(s) = \sum_{\gamma_m} |\phi_{\gamma_m}(s)|^2 u_\gamma^m$$

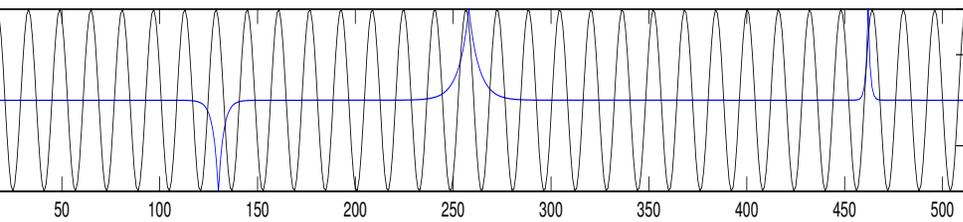
- For USMG priors with mixing RVs subband-independent and rapidly decreasing pdfs.

$$P(\Phi_m | Y_s; \mathbf{U}_\gamma^m) = \frac{\mathcal{N}(Y_s; \mu_m, \hat{V}_m + \sigma_\varepsilon^2)}{\sum_{m'=1}^M \mathcal{N}(Y_s; \mu_{m'}, \hat{V}_{m'} + \sigma_\varepsilon^2)}$$

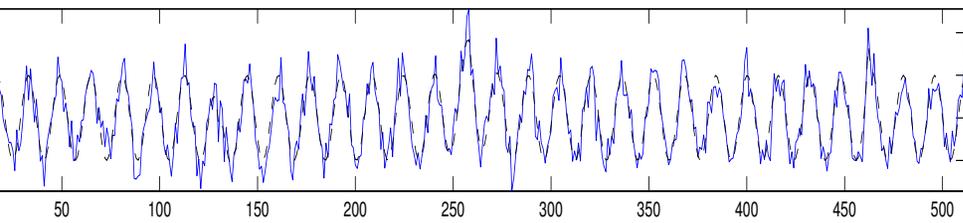
- For many usual transform bases, \hat{V}_m is exactly (or to a good approximation):

$$\hat{V}_m = (\approx) \sum_{\gamma_m} \underbrace{\hat{u}_\gamma^m}_{\text{Empirical bayes estimate}}$$

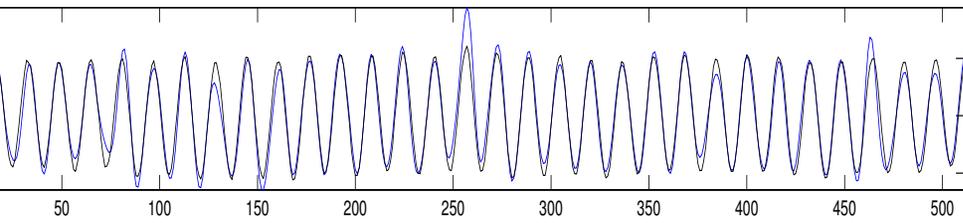
1D Examples



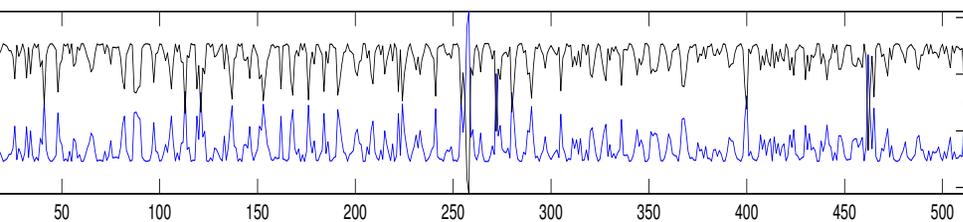
Original and noisy $\text{PSNR}_m = 15.25$



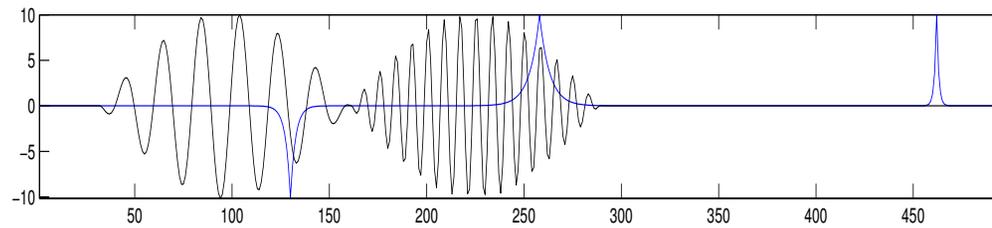
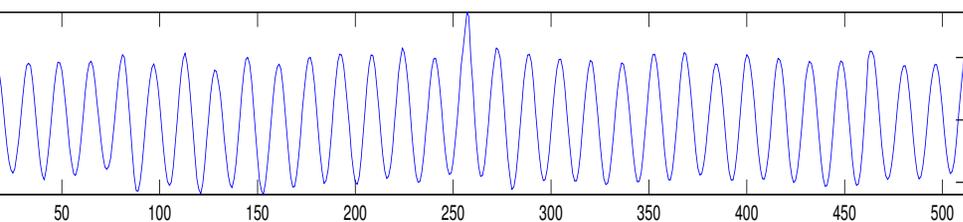
$\text{PSNR}_w = 22.27$ $\text{PSNR}_d = 22.63$



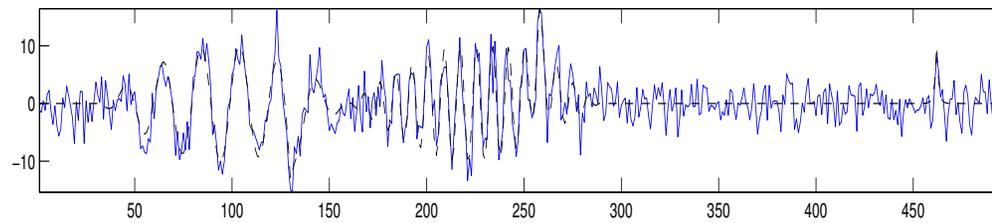
$P(\phi_m | Y)$



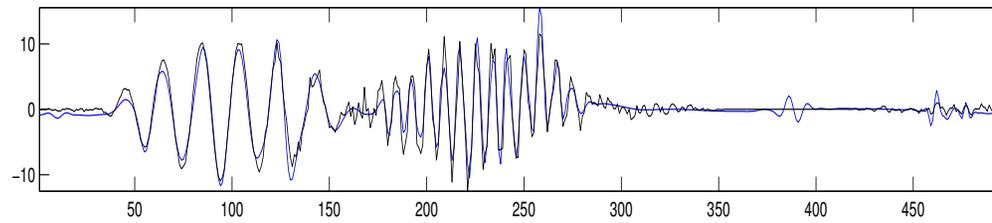
$\text{PSNR}_{cb} = 24.62$



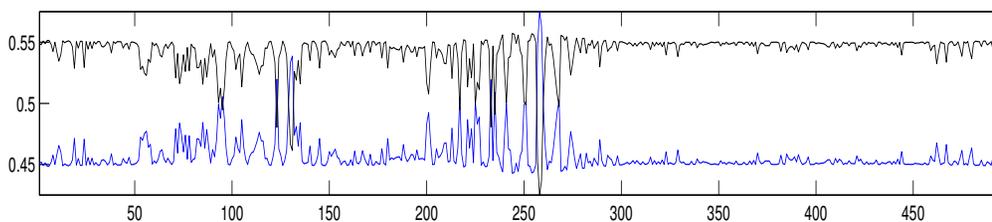
Original and noisy $\text{PSNR}_m = 15.25$



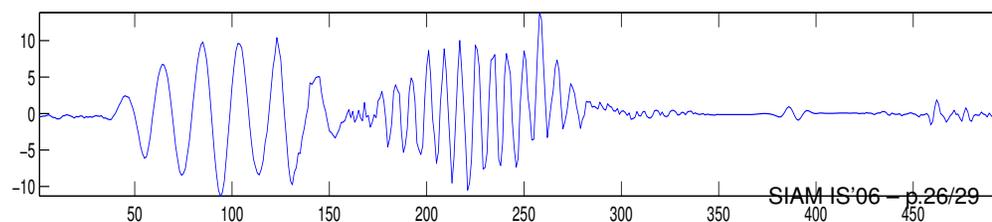
$\text{PSNR}_w = 21.76$ $\text{PSNR}_d = 23.30$



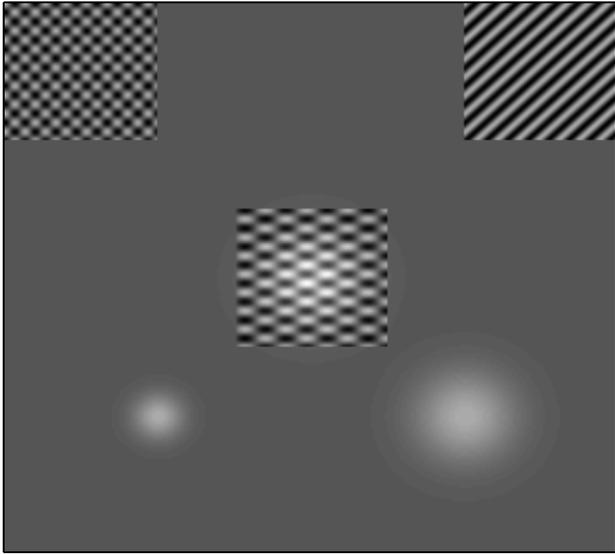
$P(\phi_m | Y)$



$\text{PSNR}_{cb} = 24.06$

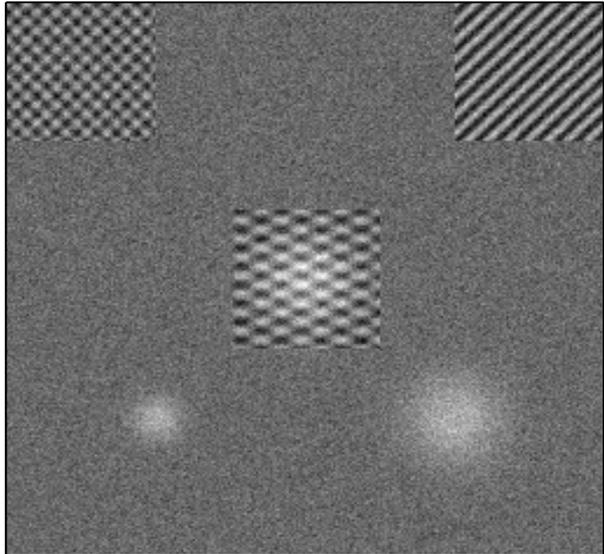


Original

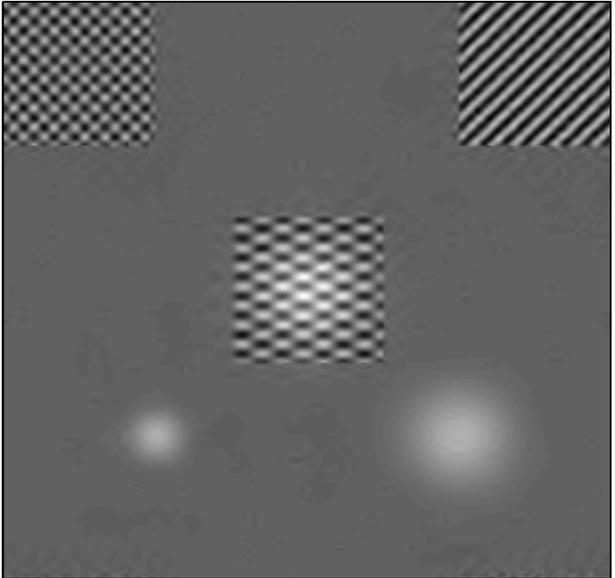


$\text{PSNR}_w = 32.95$

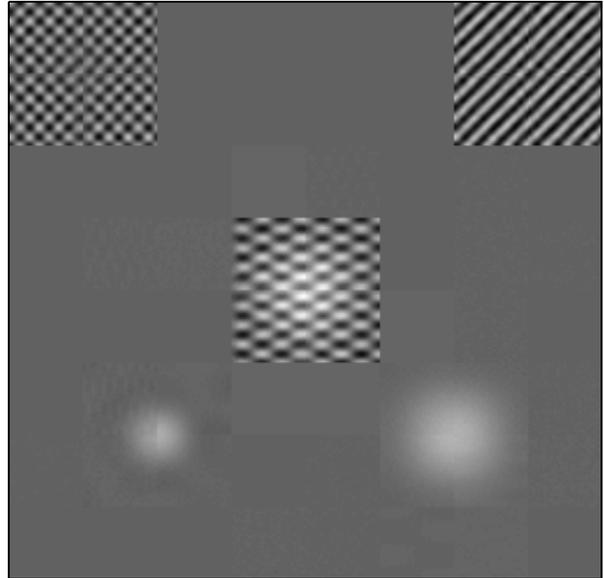
Noisy $\text{PSNR}_{in} = 20.62$



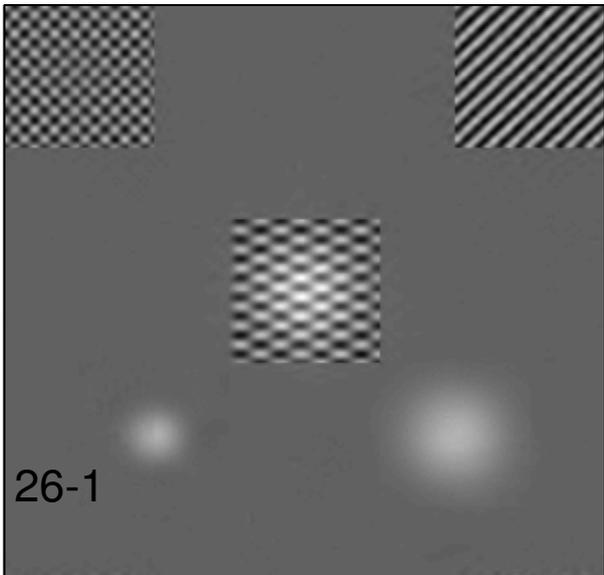
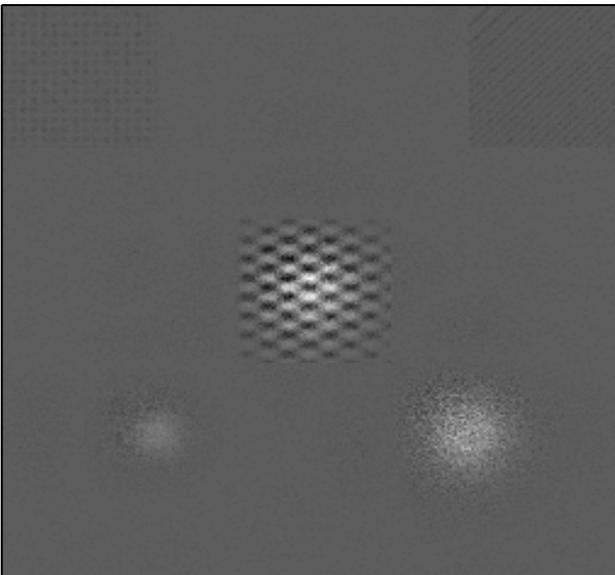
$\text{PSNR}_d = 32.39$



$P(\Phi_w | Y)$



$\text{PSNR}_{cb} = 34.47$



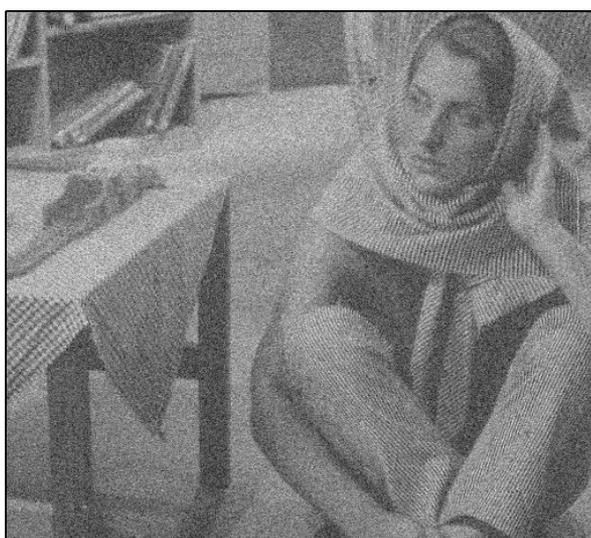
26-1

Original



$\text{PSNR}_w = 23.95$

Noisy $\text{PSNR}_{in} = 16.09$



$\text{PSNR}_d = 24.22$



$P(\Phi_w | Y)$



$\text{PSNR}_{cb} = 25$



Combined PSNR=25



MBKF UDWT PSNR=24.7



MBKF FDCT PSNR=26.12



Conclusion and perspectives

- A flexible statistical prior to model both marginal and joint statistics of sparse representation coefficients.
- Univariate and multivariate properties derived and special cases fully considered.
- Application to statistical modeling of real images.
- Bayesian term-by-term MF and MAP estimators also derived.
- Combined MF denoising with bases.
- Extension to other transforms.
- Extension to handle more rigorously dependencies (e.g. global MRF).

More details at

<http://www.greyc.ensicaen.fr/~jfadili>

**Thanks
Any questions ?**